

# Low Dose Hydroxychloroquine is Associated with Lower Mortality in COVID-19: A Meta-Analysis of 27 Studies and 44,684 Patients

**Augusto Di Castelnuovo**

Mediterranea Cardiocentro, Napoli

**Simona Costanzo**

Department of Epidemiology and Prevention, IRCCS Neuromed, Pozzilli

**Antonio Cassone**

Polo della genomica, genetica e biologia, Università di Siena

**Roberto Cuda**

Fondazione Policlinico Universitario A. Gemelli IRCCS Università Cattolica S. Cuore Roma

**Giovanni de Gaetano**

Department of Epidemiology and Prevention, IRCCS Neuromed, Pozzilli

**Licia Iacoviello** (✉ [licia.iacoviello@moli-sani.org](mailto:licia.iacoviello@moli-sani.org))


Department of Epidemiology and Prevention, IRCCS Neuromed, Pozzilli

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## Research Article

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# Abstract

**Background:** Hydroxychloroquine (HCQ) was proposed as potential treatment for COVID-19, but its association with mortality is not well characterized. We conducted two meta-analyses to evaluate the association between HCQ (with or without azithromycin (AZM)) and total mortality in COVID-19 patients.

**Methods:** Articles were retrieved until November 10<sup>th</sup>, 2020 by searching in seven databases. Data were combined using the general variance-based method.

**Results:** A total of 27 articles were found (N=44,684 patients, including N=7,803 from 5 randomized clinical trials (RCTs)). Overall, the use of HCQ was associated with 22% lower mortality risk (pooled risk ratio: 0.78, 95%CI: 0.67 to 0.91; I<sup>2</sup>=80%, random effects). This association was found reversed when pooling only RCTs (16.7% of the overall weight; pooled risk ratio: 1.11, 0.99 to 1.24) or studies in which daily dose >400 mg or total dose >4,400 mg were used (pooled risk ratio: 1.10, 95%CI: 0.99 to 1.23 in both cases). Overall, HCQ+AZM (10 studies) was also associated with 28% lower mortality risk, but uncertainty was large (95%CI: 0.48 to 1.08; P=0.11). Use of HCQ was not associated with severe adverse events.

**Conclusions:** HCQ use was not associated with mortality reduction in COVID-19 patients when 5 RCTs only were evaluated, while a 9% to 33% reduced mortality was observed when observational studies were also included. The association was mainly apparent in studies that used lower doses of HCQ. These findings can help disentangling the debate on HCQ use in COVID-19.

## Introduction

The aminoquinoline hydroxychloroquine (HCQ) is an anti-malaria drug, with immunomodulatory and anti-thrombotic properties, currently used in the treatment of autoimmune diseases like rheumatoid arthritis, systemic lupus erythematosus and anti-phospholipid syndrome [1,2]. At the beginning of the pandemic, it was proposed as a possible therapy in COVID-19 patients since it could directly inhibit viral entry and spread in several *in vitro* and *in vivo* models [3]. Indeed, HCQ has been used in Ebola virus disease [4], human immunodeficiency virus (HIV) infection [5], SARS-CoV-1 infection and the Middle East Respiratory Syndrome [6].

Despite the lack of evidence of efficacy from few randomized clinical trials, HCQ became very popular and widely used by many clinicians. In Italy over 70% of COVID-19 hospitalised patients were treated with HCQ [7]. On the other hand, this drug was given an inexplicable political connotation that shifted the focus more to a political battle than to a discussion based on scientific evidence. The publication of a very questionable study [8] by one of the most reliable scientific journals showing that the use of HCQ was associated to an increased risk of death, led to the stop of several clinical trials in their tracks and of the HCQ arm in the Recovery trial [9]. This Lancet [study was retracted](#) 13 days after publication [10], because its data turned out to be fabricated. Anyway, several national agencies for drug regulation decided to suspend the authorization to use HCQ for COVID-19. Consequently, the review process of papers on HCQ became increasingly difficult, creating a publication bias that affected all meta-analyses published until now.

However, a number of questions remain open on the relationship between HCQ treatments in COVID-19 patients: is there a dose issue? Does mortality rate of a population or the severity of the disease affect HCQ efficacy? Is there any interaction with other anti-COVID19 drugs? In the last few months, at least three large, well-conducted observational studies have been published showing a protective effect of HCQ on mortality risk in hospitalized COVID-19 patients [7,11,12]. All three studies have not been included in previous meta-analysis [13], and used HCQ doses lower than those administered in randomized clinical trial (RCT), such as the Recovery or the Solidarity trials [9,14].

Therefore, we decided to conduct an updated meta-analysis on observational and RCT studies on HCQ use and the mortality outcome in patients hospitalized for COVID-19. We also performed subgroup analyses to dissect whether treatment effects differ according to methodological or clinical characteristics of the primary studies.

## Methods

This study was conducted according to the recommendations outlined in the Cochrane Handbook for Systematic Reviews of Interventions, version 5.1.0, and reported in line with the PRISMA statement. Institutional review board approval was not required as the study did not directly involve human participants.

### *Search strategy*

Flow diagram for study selection is reported in Figure 1. Articles published in English were retrieved from inception to November 10<sup>th</sup>, 2020 by searching in Medline, Embase, PubMed, Web of Science, Cochrane Central Database, MedRxiv and Preprints.org, with the search terms: “(COVID-19 OR Cov-Sars-2) AND (hydroxychloroquine OR chloroquine)”. In addition, the reference lists of relevant articles for potential studies were also manually reviewed. After initial search, the duplicate results were removed. The remaining articles were screened for relevance by their titles and abstracts by two of us independently (SC and ADC). All selected potential articles were then reviewed by the remaining investigators to ensure their eligibility for inclusion. Disagreements about eligibility of the literature were resolved by consensus based on the agreements of all investigators.

To be included in this meta-analysis, the study had to meet the following criteria: (1) clinical trials or cross-sectional studies or cohort studies; (2) quantitatively investigating the difference in mortality risk in unselected COVID-19 patients according to use or not of HCQ.

Twenty-eight articles were identified [7,9,11,12,14-37]. For ten of them [11,15,17,19,21,23,25,31,32,34] it was possible to extract data necessary for comparing HCQ+AZM *versus* no HCQ+AZM. For the other studies, it was not possible to systematically distinguish if HCQ therapy was complemented or not with AZM. The RCT by Mitjà et al. [29] was not included since the Authors found zero deaths both in HCQ and control group. The study of Mehra et al. [8] was excluded since the Authors retracted it owing to several concerns on the veracity of the data [10].

### *Data Extraction and data analysis*

For each study, odds ratio (OR) or hazard ratio (HR) and/or number of events (number of deaths and number of total COVID-19 patients) in both the HCQ (or HCQ+AZM) and respective control groups were extracted. If available, measure of association adjusted for covariates were retrieved. Number of events were used to calculate relative risk and 95% confidence intervals (CIs) when other measures of association were not available from the primary study. The following information was also extracted: study design, if the article was not peer-reviewed, region, level of adjustment, sample size, mortality rate in the entire cohort, percentage of patients treated with HCQ, mean duration of the treatment, mean daily dose after the first day and mean total dose of HCQ used. The total dose of HCQ was calculated as the sum of the amount of drug used in the first day plus daily dose multiplied by number of days of treatment after the first.

### **Pre-specified subgroup analyses have been conducted for all the additional characteristics retrieved.**

All analyses were performed using standard statistical procedures provided in RevMan5.1 (The Cochrane Collaboration, Oxford, United Kingdom). Data were combined using the general variance-based method, that requires information on the relative risk (or OR or HR) estimate and their 95% CI for each study. 95% CI were used to assess the variance and the relative weight of each study. Heterogeneity was assessed using the Higgin's  $I^2$  metric. Fixed and random effects were

considered, but due to the large heterogeneity observed, findings from random effects were considered as primary analysis. The hypothesis that publication bias might have affected the validity of the estimates was visually tested by a funnel plot-based approach.

## Results

### *Characteristics of the studies*

The workflow of the process of study selection is reported in Figure 1. A total of 28 articles were found in the search. Twenty-seven of these eligible studies were enrolled for analysing the association with mortality of HCQ use in patients with COVID-19, and 10 of them were valuable for analysing the association of HCQ+AZM.

The main characteristics of the 27 studies included in the meta-analysis are shown in Table 1. We found 5 RCT studies [9,14,17,33,35] and 22 observational studies; 5 articles were not published in peer reviewed journals; 3 studies reported not adjusted measure of association between HCQ and mortality; 13 studies have been conducted in Europe, 10 in North America and 4 in other countries. The outcome considered was total mortality, with the exception of Geleris et al. [18] in which the authors used a combined endpoint formed by intubation or death; for the large majority of the studies, the mortality was intra-hospital. In all studies the control group was formed by patients without HCQ exposure (HCQ or HCQ+AZM). All studies included adult men and women COVID-19 patients, with the exception of two RCT [9,35], that included a portion of individuals with uncertain positivity to Sars-CoV-2. A total of N=44,684 COVID-19 patients (including N=7,803 from the 5 RCTs) were counted in the meta-analysis of HCQ, and N=1,821 in the meta-analysis of HCQ+AZM.

### *Pooled analysis, HCQ as exposure*

Forest plot on the association between HCQ and mortality is reported in Figure 2. Pooling of data from 22 observational studies, which accounted for 83.3% of the total weight, the use of HCQ has been associated with 28% lower mortality risk (pooled risk ratio: 0.72, 95%CI: 0.61 to 0.85; high level of heterogeneity:  $I^2=77%$ , random effects). Conversely, the association was found reversed in the 5 RCTs (16.7% of the weight; pooled risk ratio: 1.11, 95%CI: 0.99 to 1.24;  $I^2=0%$ ). Grand total by pooling of both RCTs and observational studies (weight 100%) pointed to a relative reduction of mortality in support of HCQ by 22% (9% to 33%) (figure 2). After the exclusion of the study by Geleris et al. [18] (which used a combined outcome of intubation and death) findings remained unchanged (pooled risk ratio: 0.76, 95%CI: 0.65 to 0.90;  $I^2=80%$ ).

Subgroup analyses according to main features of primary studies are presented in Table 2. The association of HCQ with lower mortality was observed with very low differences in all subgroups, with the exception of dose grouping.

The reduced mortality was in fact confined to studies that used a daily dose  $\leq 400$  mg (as estimated in days of treatment after the first, in which a higher (double for most) dose of drug was administered); pooling n=9 studies which used more than 400 mg of HCQ daily resulted in an overall measure of association equal to 1.10 (95%CI: 0.99 to 1.23; table 2 and supplementary figure 1). Also, the magnitude of the negative association with mortality was higher in studies which used HCQ for 5 or less days (table 2 and supplementary figure 2). Pooling studies which used more than 4,400 mg of HCQ during the entire phase of treatment set an overall measure of association equal to 1.10 (95%CI: 0.99 to 1.23; table 2 and supplementary figure 3). These findings were confirmed in subgroup analyses restricted to observational studies: the pooled measure of association was equal to 0.98 (95%CI: 0.54 to 1.77) and 1.05 (95%CI: 0.73 to 1.53) in studies with a total HCQ dose  $>4,400$  mg (n=3) or a daily HCQ dose  $>400$  mg (n=5), respectively.

Subgroup analysis according to both dosing ( $\leq 400$  mg/d or  $>400$  mg/d) and duration of treatment ( $\leq 5$  days or  $>5$  days) confirms the modification of effect by overall dosing of HCQ, with a maximum protection observed in studies with  $\leq 400$  mg/d and  $\leq 5$  days (0.63, 95%CI: 0.46 to 0.86; supplementary figure 4) as opposed to studies with  $>400$  mg/d and  $>5$  days (1.10, 95%CI: 0.99 to 1.23; supplementary figure 4). Interestingly, pool of studies with both low doses and long duration also provided an overall association against mortality (0.74, 95%CI: 0.55 to 1.00; supplementary figure 4).

### ***Pooled analysis, HCQ+AZM as exposure***

Figure 3 reported random forest for 10 studies comparing HCQ+AZM. Use of the combination HCQ+AZM was associated with 28% lower mortality risk, with very large uncertainty (pooled risk ratio: 0.72, 95%CI: 0.48 to 1.08; P for testing of overall effect=0.11; high level of heterogeneity:  $I^2=91\%$ , random effects).

By visual inspection of funnel plots (Supplementary Figure 5), we failed to observe any selection bias for both meta-analyses.

### ***Pooled analysis of adverse events***

Pooled analyses of RCTs on the relationship between HCQ and incidence of adverse effects are reported in supplementary figure 6. HCQ use was associated with an increased risk of adverse effects of any type (panel A). On the contrary, patients treated with HCQ in RCTs showed a similar rate of serious adverse events (panel B), as that non-treated with HCQ (pooled risk ratio: 1.17, 95%CI: 0.89 to 1.53; P for testing of overall effect=0.27).

## **Discussion**

In a meta-analysis of 27 studies (5 RCTs) involving 44,684 COVID-19 patients, the use of HCQ was associated with a 22% lower risk of total mortality. The association was apparent by pooling 22 observational studies and was more evident in studies which used lower overall HCQ dosing. No association was found pooling 5 RCTs. Use of HCQ was not associated with severe adverse events.

The potential for selection bias in observational studies is an important issue. The decision from the clinician to utilize or not a drug may depend on comorbidities and baseline risk of the patient. However, in the pandemic, and in absence of guidelines and specific anti COVID-19 drugs, allocation of HCQ in observational studies was not associated systematically with a lower or higher baseline risk profile. For example, in the CORIST study [7] patients receiving HCQ were more likely younger and less likely had ischemic heart disease, cancer or chronic kidney disease, but, on the contrary, they were more likely men and had higher levels of C-reactive protein. As a consequence, it is not clear if in that particular study HCQ patients were potentially at higher or lower risk of a negative prognosis. In attempting to account for baseline differences between patients who received HCQ and those who did not, we used the results for adjusted measure of association for each study, and this was possible for 22 out of 27 studies included in the meta-analysis. After the exclusion of 3 unadjusted studies [21,24,35], the strength of the overall association of HCQ with mortality unchanged. Although we attempted to control for potential confounding factors inherent to patient and clinical characteristics, it is possible that unmeasurable confounding still remains, and this may explain the different finding between observational and RCT studies. However, it is hard to determine which are, if any, the unmeasured characteristics that have confused so strongly the association between HCQ and mortality in COVID-19 that was observed in observational studies. In fact, these features must be a) unmeasured in observational studies; b) associated with mortality in COVID-19 and c) associated with HCQ use, in a way that when the risky conditions are present the clinicians tend systematically to avoid using HCQ. For example, HCQ is contraindicated in patients with

cardiomyopathy but this condition has been mostly measured in observational studies and was not recognised as a risk factor for mortality in COVID-19 patients.

The dissimilar findings between observational and RCTs we found might also be explained by differences in HCQ dosage [38]. Interestingly, we observed that the reduced mortality associated with HCQ treatment is actually confined to studies that used a daily dose  $\leq 400$  mg, or a total dose  $\leq 4,400$  mg or which used HCQ for 5 or less days. Obviously these three conditions largely overlapped in studies, that we can now designate as “at low HCQ dosing”. Remarkably, 4 over 5 RCTs are in this category. In detail, the RECOVERY [9] and the SOLIDARITY study [14] used 800 mg/day for 9 or 10 days (after the first), respectively and a total dose of 9200 or 10000 mg of HCQ (including the dose at first day) respectively, a very high dose regimen as compared to the rest of studies, particularly of the observational ones.

The possibility that HCQ reduced the risk of negative prognosis in COVID-19 patients when only administered at “low dose” cannot be here undoubtedly proven starting from our findings, but it is a plausible hypothesis that may explain the different result between observational and RCT studies and, more importantly, might be useful in disentangling the debate on HCQ use in COVID-19. If our hypothesis of “low doses-short duration” is correct, it follows that immunomodulation and inflammation occurs quite early after infection with SARS-CoV-2, as also suggested by the benefit of HCQ treatment in patients with higher CRP [7]. In this line, it is of interest that group of studies with low doses and long duration also provided an overall association against mortality. This finding would also support the immunomodulatory hypothesis as potential mechanism of HCQ action [37,39] as it implies a cytokine rebound when the treatment is stopped; if this happens at a critical moment, it could worsen the patient condition, thus vanishing the (potential) effect of HCQ [40].

High levels of HCQ administration were used in RCTs to maximise the antiviral activity of the drug that was considered to be the main mechanism of action of HCQ in this context. In some studies, the inverse association of HCQ with inpatient mortality was more evident in elderly, in patients who experienced a higher degree of COVID-19 severity or having elevated C-reactive protein levels [7], suggesting that the anti-inflammatory potential of HCQ may have had a more important role than its antiviral properties. HCQ, indeed, beside an antiviral activity, may have both anti-inflammatory and anti-thrombotic effects [3]. This can justify its effect in reducing mortality risk, since Sars-Cov-2 can induce pulmonary microthrombi and coagulopathy, that are a possible cause of its severity [41, 42] and the lack in preventing SARS-CoV-2 infection after exposure [43]. On the other end, National guidelines suggested to use HCQ 200 mg twice daily for 5-7 days probably to maintain a better risk benefit profile hypothesizing that low doses could be more effective and safer. Indeed, non-sigmoidal, bell-shaped dose-response curves are possible with drugs having complex biological effects, multiple-binding sites or cellular and organ targets. On the other hand, anti SARS-2-CoV-2 activity of HCQ has been confirmed in Vero cells [44]. HCQ is also reported to reduce secretion of IFN- $\gamma$  and IL-17 in activated Th1 and Th17 cells, respectively [45].

The concomitant use of azithromycin seems to not neither increase nor decrease the effect, if any, of the HCQ since the combination of the two drugs was associated with a lower mortality risk at very similar extent to that observed for HCQ alone, but the assumption is inconclusive because of the very large uncertainty in the findings.

A main concern with HCQ treatment have been its side effects, in particular a severe cardiovascular toxicity. Indeed, HCQ can cause prolongation of the QT interval on electrocardiogram [46], which could be exacerbated by coadministration with azithromycin, widely prescribed as co-treatment in Covid-19 treatment. Our meta-analysis of data from RCTs, that allowed a proper evaluation of side effects, shows that use of HCQ was associated with an increase in side effects of any type, but not of major type, including cardiovascular events. This despite the high prevalence of cardiovascular disease in patients with COVID-19 or the high dose used in RTCs.

This meta-analysis has the strength of including all available data recently published and that have not been included in previous meta-analyses [13,47], and of considering modification of effect by dosing of HCQ. As major limitations, we recognise that the majority of the primary studies were observational, the pooled findings suffer of a high degree of heterogeneity and that results in observational and RCT studies were different.

## Conclusions

In conclusion, HCQ was not associated with decreased mortality in COVID-19 patients when only RCTs were pooled, but it was associated with 9% to 33% reduced mortality when observational studies were also included. The association was mainly apparent by pooling studies using lower doses of HCQ. Use of HCQ was not associated with severe adverse events.

These results should be considered with caution, because the majority of the studies included were observational and retrospective and the possibility of confounding could not be fully excluded. However, at present, this is the largest comprehensive quantitative overview on the association of HCQ with mortality in COVID-19 patients, and our findings underscoring HCQ dosage effects can help disentangling the debate on HCQ use and encourage the planning of RCTs using low doses of HCQ (not necessarily with a short duration of the treatment) in hospitalised COVID-19 patients.

## List Of Abbreviations

RCT mean randomized controlled trial; RR means relative risk; OR means odds ratio; HR means hazard ratio; 95%CI means 95% confidence interval; SE means standard error; HCQ means hydroxychloroquine; AZM means azithromycin

## Declarations

### Ethics approval and consent to participate

Not applicable

### Consent for publication

Not applicable

### Availability of data and materials

All data generated or analysed during this study are included in this published article [and its supplementary information files] or in the primary published studies included in the meta-analysis.

### Competing interests

The Authors report no conflict of interest related to the current work.

### Funding

None

### Authors' contributions

ADC and LI contributed to the conception and design of the work and interpretation of data; SC and ADC managed study selection and data extraction and critically reviewed the results; SC analysed the data; ADC and LI wrote the

paper; LI, AC, RC and GdG originally inspired the research and critically reviewed the manuscript.

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## Author contributions

ADC and LI contributed to the conception and design of the work and interpretation of data; SC and ADC managed study selection and data extraction and critically reviewed the results; SC analysed the data; ADC and LI wrote the paper; LI, AC, RC and GdG originally inspired the research and critically reviewed the manuscript.

All Authors approved the final version of the manuscript.

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## Tables

**Table 1** Characteristics of the studies included in the meta-analysis

Study	Country	Type of study	N. of patients	HCQ treatment (%)	Mortality (%)	Duration (day, median)	Daily dose (mg/day)	Total dose (mg)
Albani F <sup>15</sup>	Italy	retrospective	816	25.9	25.7	6	400	2400
Arshad S <sup>11</sup>	USA	retrospective	1611	74.6	18.1	5	400	2400
Ayerbe L <sup>16</sup>	Spain	retrospective	2019	92	14.5	5	400	2800
Catteau L <sup>12</sup>	Belgium	retrospective	8075	56.2	21.8	5	400	2400
Cavalcanti AB <sup>17</sup>	Brazil	RTC	332	47.9	2.6	7	800	5600
Di Castelnuovo A <sup>7</sup>	Italy	retrospective	3451	76.3	16.7	10	400	4400
Geleris J <sup>18</sup>	USA	retrospective	1376	58.9	16.8	5	400	2800
Ip A <sup>19</sup>	USA	retrospective	2256	84.8	22	5	400	2400
Horby P <sup>9</sup>	United Kingdom	RTC	4716	33.1	25.6	10	800	9200
Kalligeros M <sup>20</sup>	USA	retrospective	108	33.3	0.9	5	NA	NA
Lagier JC <sup>21</sup>	France	retrospective	400	25.3	0.9	10	600	6000
Lammers AJJ <sup>22</sup>	The Netherlands	retrospective	689	27.4	18	NA	NA	NA
Lauriola M <sup>23</sup>	Italy	retrospective	80	21.3	38.7	10	600	6000
Lecronier M <sup>24</sup>	France	retrospective	80	47.5	31	NA	400	NA
Magagnoli J <sup>25</sup>	USA	retrospective	277	51.1	17.3	5	400	2000
Mahévas M <sup>26</sup>	France	retrospective	173	50.8	9.4	2	600	1200
Membrillo FJ <sup>27</sup>	Spain	retrospective	166	73.5	28.9	5	400	2800
Mikami T <sup>28</sup>	USA	retrospective	2820	73.7	21.7	5	NA	NA
Mitjà O <sup>29*</sup>	Spain	RTC	293	46.4	0	7	400	3600
Paccoud O <sup>30</sup>	France	retrospective	84	45.2	6.2	10	600	6000
Pan H <sup>14</sup>	30 Countries	RTC	1853	51.1	11.1	11	800	10000

Worldwide								
Rosenberg ES <sup>31</sup>	USA	retrospective	492	55.1	20.3	NA	800	NA
Sbidian E <sup>32</sup>	France	retrospective	4415	14.1	21.4	10	400	4200
Self WH <sup>33</sup>	USA	RTC	479	50.5	10.4	5	400	2400
Singh S <sup>34</sup>	USA	retrospective	1402	50	11.7	NA	NA	NA
Skipper CP <sup>35</sup>	United States and Canada	RTC	423	50	0.5	5	600	3800
Sulaiman T <sup>36</sup>	Saudi Arabia	retrospective	5541	32.8	1.1	5	400	2400
Yu B <sup>37</sup>	China	retrospective	550	8.7	44.9	8	400	3200

RTC means randomized clinical trial; NA means not available. \*Since in this study have been observed zero deaths both in HCQ and control group, it was included in the meta-analyses of adverse effect only

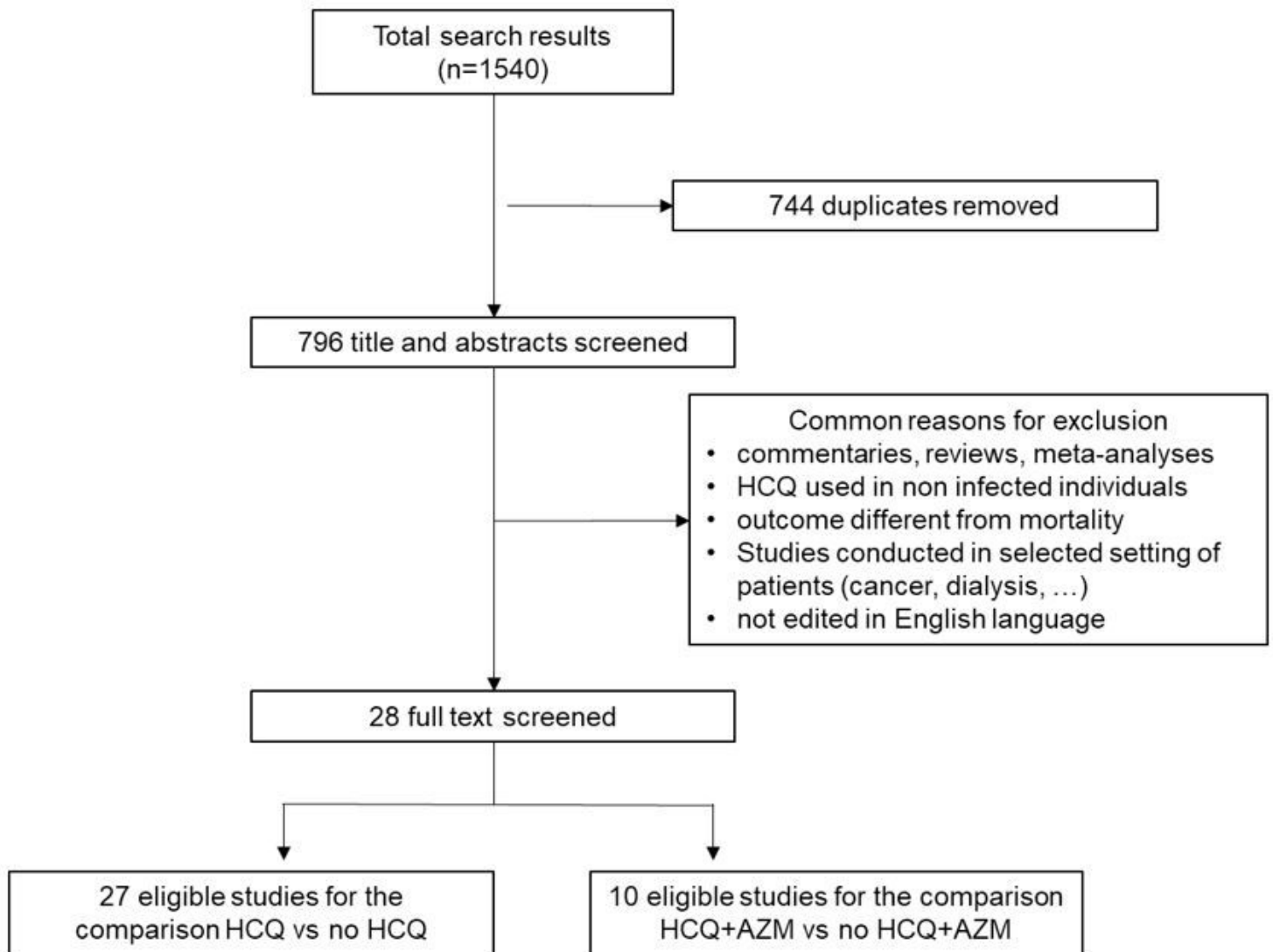
**Table 2** Pooled analysis in different subgroups of studies

Subgroups	No of Studies	I <sup>2</sup>	Pooled RR [95% CI]	
			Random effect	P value*
<b>All studies</b>	27	80%	0.78 [0.67, 0.91]	
<b>Study design</b>				<.0001
No RCT	22	77%	0.72 [0.61, 0.85]	
RCT	5	0%	1.11 [0.99, 1.24]	
<b>Peer-reviewed Studies</b>				0.69
No	5	76%	0.83 [0.57, 1.20]	
Yes	22	80%	0.76 [0.64, 0.91]	
<b>Adjustment</b>				0.54
No	3	0%	0.64 [0.35, 1.19]	
Yes	24	82%	0.78 [0.67, 0.92]	
<b>Sample size</b>				0.47
<1,000 COVID-19 patients	15	27%	0.84 [0.66, 1.06]	
≥1,000 COVID-19 patients	12	90%	0.75 [0.61, 0.91]	
<b>Mortality incidence in the sample</b>				0.94
≤20%	16	77%	0.79 [0.61, 1.01]	
>20%	11	85%	0.78 [0.63, 0.97]	
<b>HCQ treatment in the sample</b>				0.02
≤33.3%	9	59%	0.86 [0.68, 1.08]	
33.3-66.6%	12	59%	0.94 [0.78, 1.15]	
>66.6%	6	89%	0.52 [0.35, 0.75]	
<b>HCQ treatment, duration</b>				0.06
≤ 5 days	13	83%	0.65 [0.50, 0.85]	
> 5 days	10	69%	0.90 [0.73, 1.11]	
<b>HCQ treatment, total dose</b>				0.001
≤2,400 mg	8	83%	0.70 [0.52, 0.94]	
2,400 - 4,400 mg	7	81%	0.64 [0.45, 0.92]	
>4,400 mg	6	0%	1.10 [0.99, 1.23]	
<b>HCQ treatment, daily dose</b>				<.0001
≤400 mg	14	83%	0.67 [0.55, 0.82]	
>400 mg	9	0%	1.10 [0.99, 1.23]	
<b>Country</b>				0.91

Europe	13	79%	0.78 [0.64, 0.95]
North America	10	84%	0.81 [0.60, 1.11]
Others	4	81%	0.68 [0.32, 1.47]

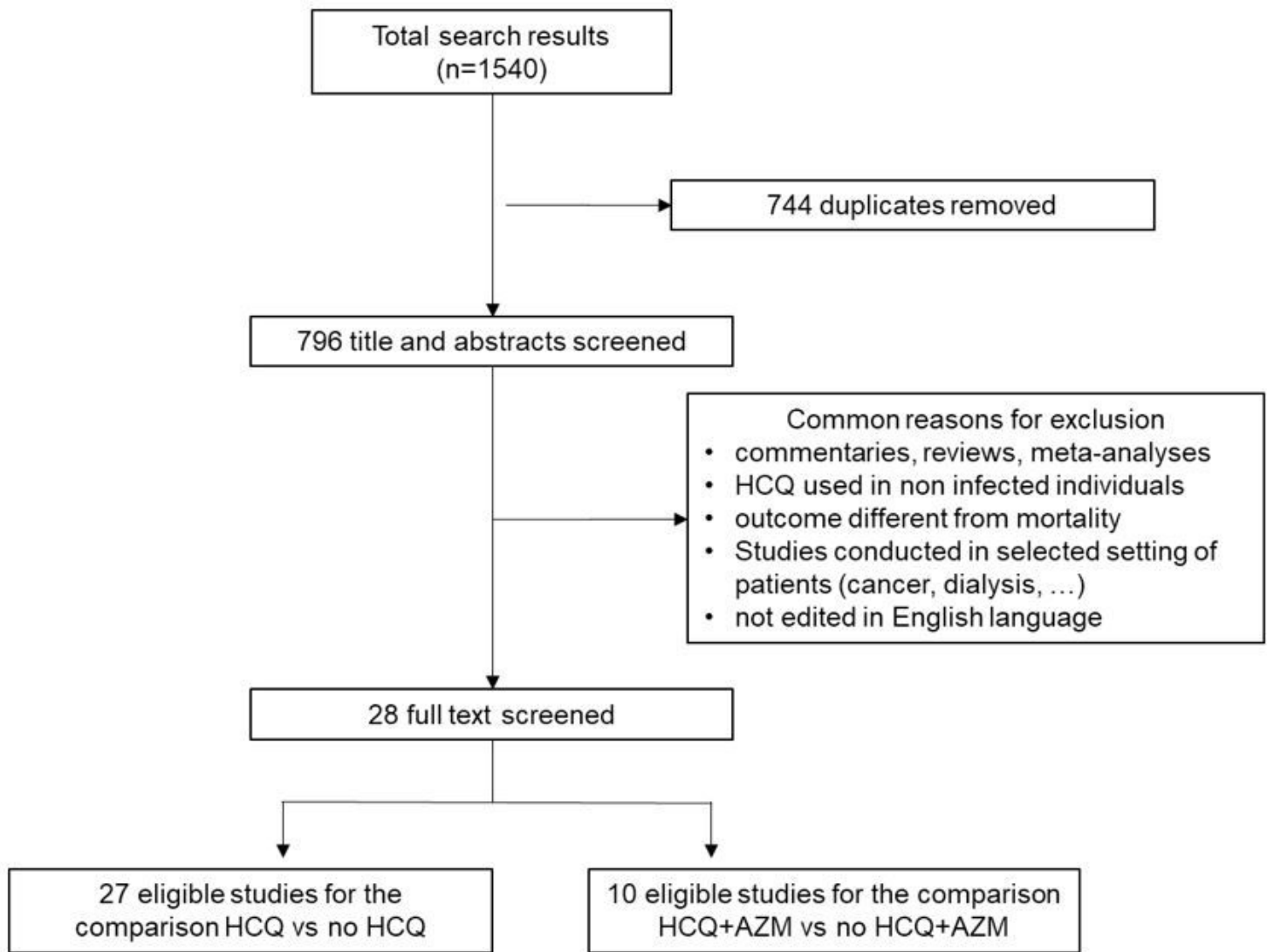
\*for difference among subgroups; RCT means randomized clinical trial; HCQ means hydroxychloroquine

## Figures



**Figure 1**

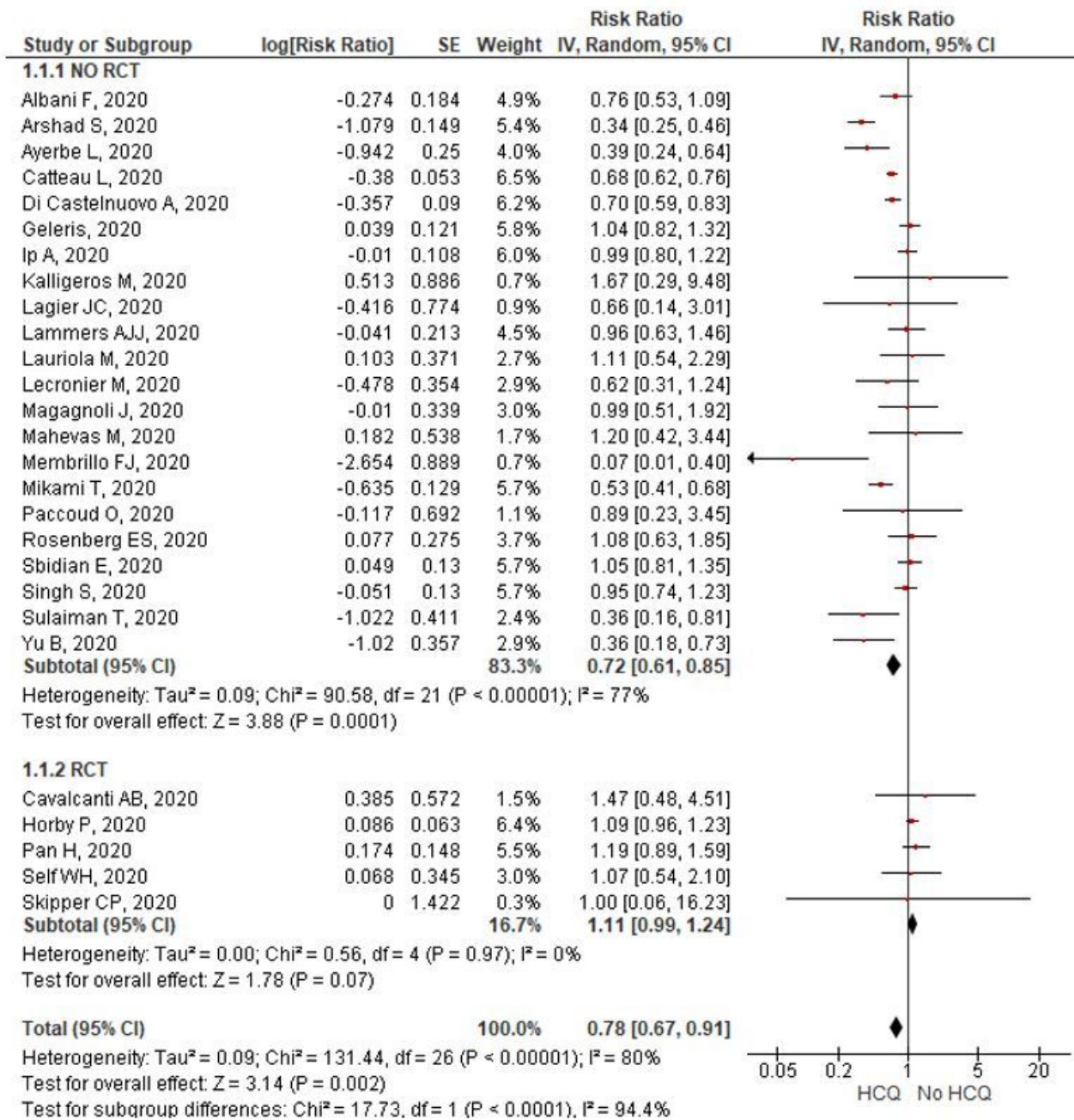
Flow diagram for study selection HCQ means hydroxychloroquine; AZM means azithromycin



**Figure 1**

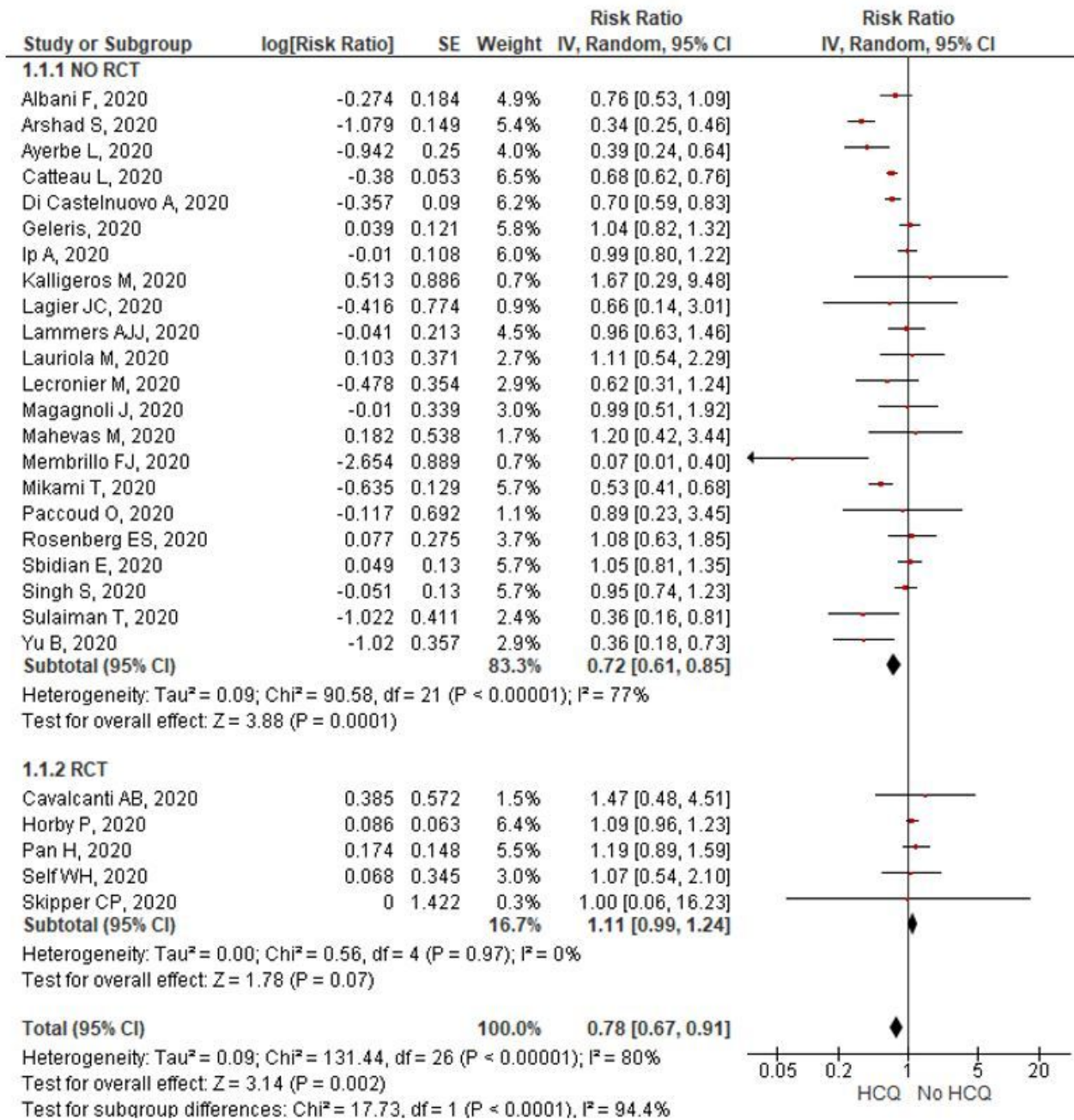
Flow diagram for study selection HCQ means hydroxychloroquine; AZM means azithromycin





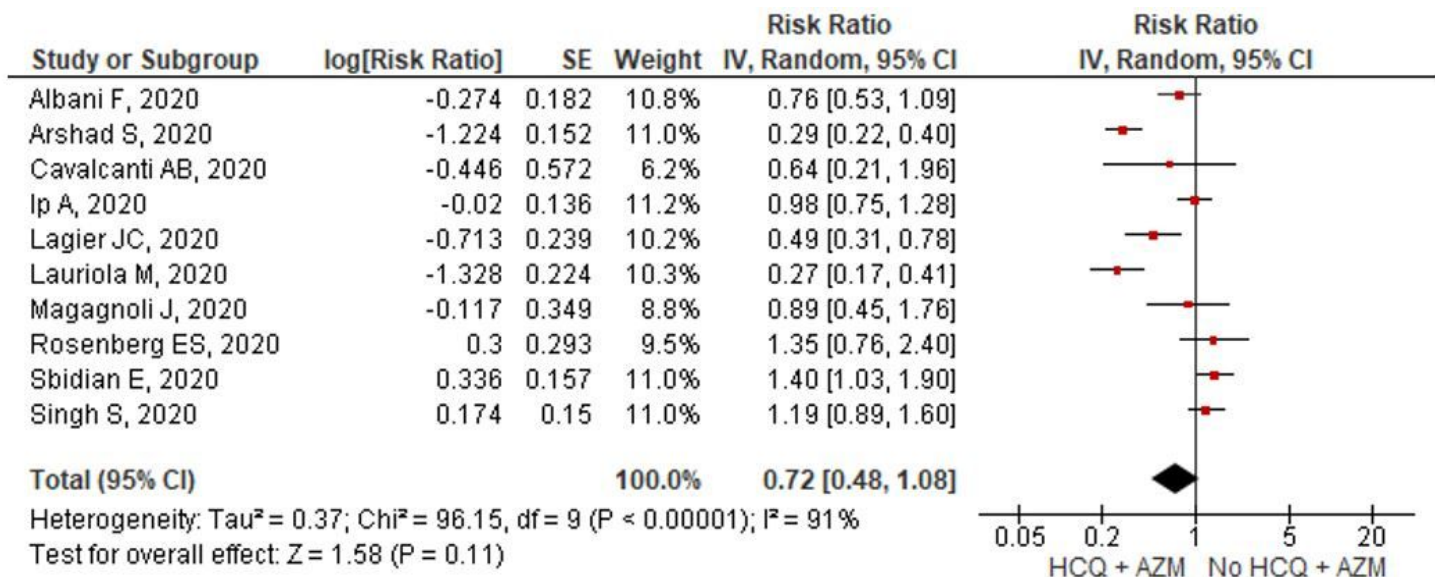
**Figure 2**

Forest plot for association of hydroxychloroquine use with COVID-19 mortality (random effects) RCT means randomized clinical trial; SE means standard error; HCQ means hydroxychloroquine



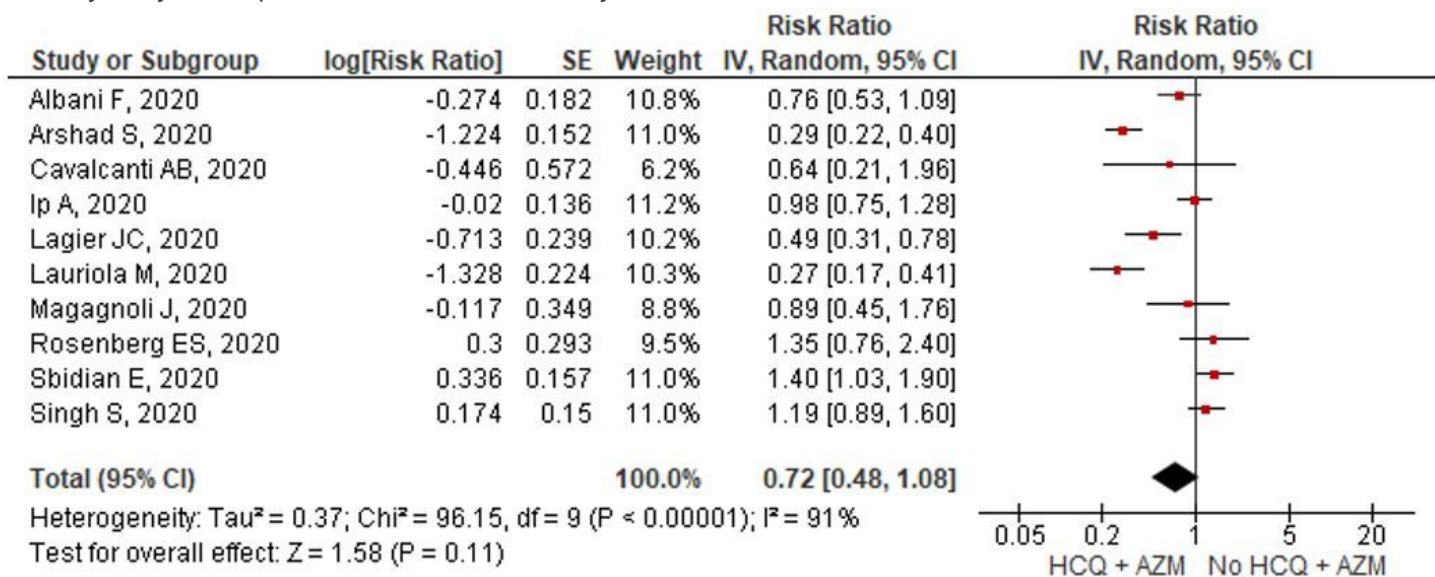
**Figure 2**

Forest plot for association of hydroxychloroquine use with COVID-19 mortality (random effects) RCT means randomized clinical trial; SE means standard error; HCQ means hydroxychloroquine



**Figure 3**

Forest plot for association of hydroxychloroquine + azithromycin use with COVID-19 mortality (random effects) HCQ means hydroxychloroquine; AZM means azithromycin



**Figure 3**

Forest plot for association of hydroxychloroquine + azithromycin use with COVID-19 mortality (random effects) HCQ means hydroxychloroquine; AZM means azithromycin

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